

# Diffusion Approximations for Counting Number of Primes

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Abstract

Quite deterministic nature of prime numbers, due to the complexity of the recurrent generating algorithms, is mimicking ‘randomness’ and stimulates to apply some of probabilistic instruments to analyze number-theoretic problems. The key issue in the probabilistic analysis in a number-theoretic framework remains an enigmatic connection between deterministic nature of integer sequences related to prime numbers and their apparent complicated (‘unpredictable’ or ‘chaotic’) behavior interpreted as ‘randomness’. We derive multiplicative and additive models with recurrent equations for generating sequences of prime numbers based on the reduced Sieve of Eratosthenes Algorithm and analyze their asymptotic behavior with the help of Riemann Zeta probability distribution. This allows to interpret such sequences as realizations of random walks on set  $\mathbb{N}$  of natural numbers and on multiplicative semigroups  $S(\mathbb{P})$  generated by set of prime numbers  $\mathbb{P}$ , representing paths of stochastic dynamical systems. We analyze in this work an additive continuous-time probabilistic model of counting function of primes  $\pi(n)$  in terms of diffusion approximation of non-Markov random walks. We assume that ‘updating’ terms  $\eta$  in the recurrent equation  $\pi(n(k+1)) - \pi(n(k)) = \eta(n(k+1))$  follow Zeta probability distribution and calculate infinitesimal characteristics of the random walk, which approximate coefficients of the corresponding stochastic differential equation. Computer modeling illustrates graphically an impressive fitting of trajectories for the original counting function, the calculated trend function, and the Brownian approximation.

“Using randomness to study certainty may seem somewhat surprising  
It is, however, one of the deepest contributions of our century to  
mathematics in general and to the theory of numbers in particular.”  
(Gérald Tenenbaum, Michel Mendès France, *The Prime Numbers and  
Their Distribution*. AMS, 2000)

In this paper we consider the sequence  $\{\pi(n)\}_{n \in \mathbb{N}}$  as a realization of a random walks  
 $\{\pi(n, \omega)\}_{n \in \mathbb{W}}$  generated by the recurrent equation

$$\pi(n_{k+1}) - \pi(n_k) = \eta(n_{k+1}) \quad \text{where} \quad \eta(n_k) = h(\min(\bar{r}(n_k))), \quad n_k = v_k(\omega) \quad (1)$$

Here  $\{v_k\}_{k \in \mathbb{N}}$  are assumed to be random variables with Zeta probability distribution.

Recall that to define a stochastic process  $\xi(t, \omega)$  with a discrete( or a continuous set  
 $\mathcal{X}$  of values we need to have a measurable space  $(\mathcal{X}, \mathcal{B})$ , where often

$\mathcal{X} \subseteq \mathbb{N}$  or  $\mathcal{X} \subseteq \mathbb{R}^d$ , a Borel  $\sigma$ -algebra  $\mathcal{B}$  of subsets on  $\mathcal{X}$ , and a set  $T$  of  
parameters  $t \in T$  such that for each  $t \in T$ ,  $\xi(t, \cdot) : \Omega \rightarrow \mathcal{X}$  is a random variable on a  
probability space  $(\Omega, \mathcal{F}, P)$ . Then, the family  $\{\xi(t, \cdot)\}_{t \in T}$  of random variables is called  
a *stochastic process* in the *phase space*  $(\mathcal{X}, \mathcal{B})$ . The parameter  $t \in T$  is usually  
interpreted as ‘discrete time’ for a countable set  $T \subseteq \mathbb{W} = \mathbb{N} \cup \{0\}$  or as ‘continuous  
time’ for the continuous interval  $T = [t_0, t_f] \subseteq \mathbb{R}^+ = [0, \infty)$ .

Then, a  $\xi(t, \omega)$  is called a *stochastic process* with a discrete or a continuous time,  
respectively. For any given elementary event  $\omega \in \Omega$ , a function  $x(\cdot) : T \rightarrow \mathcal{X}$  such  
that  $x(t) = \xi(t, \omega)$  is called a *path* (or a *trajectory*) of the random process.

Alternatively, a stochastic process can be defined as a collection of paths (*random*

elements)  $x(\cdot) = \xi(\cdot, \omega)$  in a function space  $\mathcal{X}^T = \{x(t) | t \in T\}$  where  $\omega$  that identifies each path is an elementary event in probability space  $(\Omega, \mathcal{F}, P)$ . Elements (or points)  $x \in \mathcal{X}$  are called ‘states’ of the process, and  $\xi(t, \omega)$  itself is called a process with a discrete or continuous phase space  $(\mathcal{X}, \mathcal{B})$ .

Following the historical traditions of the classical Probability Theory (and the development of Calculus, in general), we try to apply limit theorem approach to analyze behavior of infinite discrete random sequences in terms of continuous-time stochastic processes.

Let  $\vec{p}(n) = (p_1, p_2, \dots, p_k)$  be a vector of consecutive prime numbers such that

$p_1 = 2$ ,  $p_k \leq n$  and  $p_{k+1} > n$ . Index  $k$  determines here the value of function

$\pi(n) = k$  that is the number of primes less than or equal to  $n$  so that

$\vec{p}(n) = (p_1, p_2, \dots, p_{\pi(n)})$ . For each coordinate  $p_i$  of vector  $\vec{p}(n)$  we determine the

residual value  $r_i = \text{mod}(n, p_i)$ ,  $i = 1, 2, \dots, \pi(n)$ , and consider the corresponding

vector of residuals  $(r_1, r_2, \dots, r_{\pi(n)})$ . Notice that, due to the Sieve Algorithm, for an

integer  $n > 2$  to be prime it is necessary and sufficient that all coordinates

$r_i$  ( $1 \leq i \leq \sqrt{\pi(n)}$ ) of the ‘reduced’ vector of residuals  $\vec{r}(n) = (r_1, r_2, \dots, r_{\sqrt{\pi(n)}})$  be

different from zero. Meanwhile, if a random integer  $\nu$  follows Zeta distribution,

then, the events that  $\nu = n$  does not divide each of consecutive primes  $p_1, p_2, \dots, p_{\pi(\sqrt{n})}$

are independent and can be expressed as a condition:

$$\min \{r_i | 1 \leq i \leq \pi(\sqrt{n})\} > 0 \text{ or, equivalently, } \prod_{1 \leq i \leq \pi(\sqrt{n})} r_i > 0. \quad (2)$$

By using the Heaviside function  $h(x) = \begin{cases} 1 & \text{if } x > 0 \\ 0 & \text{if } x \leq 0 \end{cases}$ , we can write the recurrent

equation for  $\pi(n)$  in the form:

$$\pi(n+1) = \pi(n) + h\left(\min_{p \leq \sqrt{n}} \{\text{mod}(n, p) \mid p \in \mathbb{P}\}\right)$$

or, equivalently,

$$\pi(n+1) = \pi(n) + h\left(\min_{i \leq \sqrt{n+1}} \{r_i \mid r_i = \text{mod}(n, p_i)\}\right) = \pi(n) + h(\min(\vec{r}(n))) \quad (3)$$

which controls the occurrence of prime numbers in the sequence of all integers  $n \geq 3$ :  $h(\min(\vec{r}(n))) = 1$  if and only if  $n$  is a prime number and  $h(\min(\vec{r}(n))) = 0$  otherwise.

Consider a stochastic process approximation of non-Markov random walks  $\{\pi(n, \omega)\}_{n \in \mathbb{W}}$  such that  $\pi(n, \omega) = \pi(n)$ , with  $\pi: \mathbb{N} \times \Omega \rightarrow \mathbb{N} \cup \{0\}$  restricted to the interval of discrete ‘times’  $N_{\min} = n_0 < n_1 < \dots < n_K = N_{\max}$ :

$$\{\pi^\Delta(t_k) = \pi(n_k, \omega) \mid N_{\min} \leq n_k \leq N_{\max}\}. \quad (4)$$

Denote  $\Delta = (0 = t_0 < t_1 < \dots < t_K = 1)$  a partition of an interval  $[0, 1]$  into  $K$

subintervals, such that  $\frac{K}{\ln(N_{\max})} \rightarrow 0$  as  $N_{\max} \rightarrow \infty$ .

We can map the closed interval of real numbers  $[N_{\min}, N_{\max}] \subset \mathbb{R}$  to the interval  $[0, 1] \subset \mathbb{R}$  with an increasing continuously differentiable function  $\tau(x)$  such that  $\tau(N_{\min}) = 0$ ,  $\tau(N_{\max}) = 1$ .

In the context of our study, a suitable choice of function  $\tau$  takes the form:

$$\tau(x) = \frac{\int_{N_{\min}}^x \frac{dt}{\ln t}}{\int_{N_{\min}}^{N_{\max}} \frac{dt}{\ln t}} = \frac{Li(x) - Li(N_{\min})}{Li(N_{\max}) - Li(N_{\min})} \quad (5)$$

where  $Li(x)$  stands for the *Eulerian logarithmic integral*  $Li(x) = \int_2^x \frac{dt}{\ln t}$ .

Then,  $t_k = \tau(n_k)$  and for  $\tau^{-1}$  (the inverse of  $\tau$ ) we have  $n_k = \tau^{-1}(t_k)$  ( $k = 1, 2, \dots, K$ ).

Denote  $\Delta t_k = t_k - t_{k-1}$ . Assume that  $N_{\min} \rightarrow \infty$  and for each choice of  $N_{\min}$  a positive integer  $K$  can be taken such that  $|\Delta| = \max_{1 \leq k \leq K} \Delta t_k \rightarrow 0$ . Here a sequence of random variables  $\pi^\Delta(t_k) = \pi(n_k)$  is interpreted as a path of a walking point  $\pi^\Delta(t_k)$  that belongs to a measurable space  $(\mathcal{X}_k, \mathcal{B}_k)$  at each ‘instant of registration’  $t_k$ .

Probability distribution on the probability space  $(\Omega, \mathcal{F}, P)$  generated by the path

space  $(\mathcal{X}^\infty, \mathcal{F}^\infty) = \left( \prod_{k \in \mathbb{N}} \mathcal{X}_k, \otimes_{k \in \mathbb{N}} \mathcal{B}_k \right)$  of random walks  $\{\pi(k, \omega)\}_{k \in \mathbb{W}}$  is determined by

transition probabilities  $P\{\pi^\Delta(t_{k+1}) \in E \mid \pi^\Delta(\vec{t}_0^k) = \vec{x}_0^k\}$  where  $\vec{t}_0^k = (t_0, t_1, \dots, t_k)$ ,

$\vec{x}_0^k = (x_0, x_1, \dots, x_k) \in \mathcal{X}_0^k = \prod_{i=0}^k \mathcal{X}_i$ ,  $x_i \in \mathcal{X}_i$  ( $i=0, 1, \dots$ ),  $E \in \mathcal{B}_{k+1}$ . Existence and

uniqueness of the probability path space  $(\Omega, \mathcal{F}, P)$  follows from the theorem of

Ionescu Tulcea [41]. Notice that  $\pi : \mathbb{N} \rightarrow \mathbb{W} = \mathbb{N} \cup \{0\}$  and therefore, we set

$\mathcal{X}_k = \mathbb{W}$  for all  $k \in \mathbb{N}^*$ . To prove the weak convergence of transition probabilities for

the sequence of random walks to the diffusion process  $\hat{\pi}(t)$  on the time interval

$[0, 1]$ , consider so called *infinitesimal characteristics* of the random walks:

$$\begin{aligned} m^\Delta(t_k, \vec{x}_1^k) &= \frac{1}{\Delta t_k} E\{\Delta \pi(t_{k+1}) \mid \pi^\Delta(\vec{t}_1^k) = \vec{x}_1^k\}, \\ [\sigma^\Delta(t_k, \vec{x}_1^k)]^2 &= \frac{1}{2 \cdot \Delta t_k} E\{[\Delta \pi(t_{k+1})]^2 \mid \pi^\Delta(\vec{t}_1^k) = \vec{x}_1^k\} \\ g^\Delta(t_k, \vec{x}_1^k; \Gamma_{k+1}) &= \frac{1}{\Delta t_k} E\{I_{\Gamma_{k+1}}(\Delta \pi^\Delta(t_{k+1}) \mid \pi^\Delta(\vec{t}_1^k) = \vec{x}_1^k)\}. \end{aligned} \tag{6}$$

Here  $\Delta \pi(t_{k+1}) = \pi^\Delta(t_{k+1}) - \pi^\Delta(t_k) = \eta^\Delta(t_{k+1}) = \pi(n_{k+1}) - \pi(n_k) = \eta(n_{k+1})$

$\Delta \pi(t_{k+1}) = \pi^\Delta(t_{k+1}) - \pi^\Delta(t_k) = \pi(n_{k+1}) - \pi(n_k) = \eta^\Delta(t_{k+1}) = \eta(n_{k+1})$ ,  $\vec{t}_1^k = (t_1, t_2, \dots, t_k)$ ,

$$\pi^\Delta(\vec{t}_1^k) = (\pi^\Delta(t_1), \pi^\Delta(t_2), \dots, \pi^\Delta(t_k)) = (x_1, x_2, \dots, x_k) = \vec{x}_1^k, \quad I_\Gamma(x) = \begin{cases} 1 & \text{if } x \in \Gamma \\ 0 & \text{otherwise} \end{cases},$$

$$\Gamma_{k+1} \subset \mathcal{X}_{k+1} \setminus \{x_k\} \in \mathcal{B}_{k+1}.$$

By setting  $n_k = n_0 + k$  for all  $k = 0, 1, \dots, K$ , we have:

$$\Delta\pi^\Delta(t_{k+1}) = \pi(n_{k+1}) - \pi(n_k) = \eta(n_k + 1) = \eta^\Delta(t_{k+1})$$

$$\Delta\pi^\Delta(t_{k+1}) = \pi(n_{k+1}) - \pi(n_k) = \pi(n_k + 1) - \pi(n_k) = \eta(n_{k+1}) = \eta^\Delta(t_{k+1})$$

$$\Delta\pi^\Delta(t_{k+1}) = \pi(n_{k+1}) - \pi(n_k) = \pi(n_k + 1) - \pi(n_k) = \eta(n_k + 1) = \eta^\Delta(t_{k+1}), \text{ where}$$

$$P\{\eta^\Delta(t_{k+1}) = 1\} = r_{n_{k+1}} = \prod_{p \leq \sqrt{n_{k+1}}} \left(1 - \frac{1}{p}\right),$$

$$P\{\eta^\Delta(t_{k+1}) = 0\} = 1 - r_{n_{k+1}}$$

We have then,

$$m^\Delta(t_k, \vec{x}_k) = \frac{1}{\Delta t_k} \cdot E\{\eta(n_{k+1})\} = \frac{1}{\Delta t_k} \cdot \prod_{p \leq \sqrt{n_{k+1}}} \left(1 - \frac{1}{p}\right) \quad (7)$$

Similar, since  $\eta(n_{k+1}) = [\eta(n_{k+1})]^2$ , we have

$$[\sigma^\Delta(t_k, \vec{x}_k)]^2 = \frac{1}{\Delta t_k} \cdot E\{[\eta(n_{k+1})]^2\} = \frac{1}{\Delta t_k} \cdot \prod_{p \leq \sqrt{n_{k+1}}} \left(1 - \frac{1}{p}\right) \quad (8)$$

By applying the first Merten's theorem to (7) and (8), we have

$$m^\Delta(t_k, x_{n_k}) = \frac{1}{\Delta t_k} \cdot \frac{c}{\ln(n_k + 1)} \cdot \left[1 + O\left(\frac{1}{\ln(n_k + 1)}\right)\right] \quad (9)$$

$$[\sigma^\Delta(t_k, x_{n_k})]^2 = \frac{1}{\Delta t_k} \cdot \frac{c}{\ln(n_k + 1)} \cdot \left[1 + O\left(\frac{1}{\ln(n_k + 1)}\right)\right]$$

### Lemma 1.

For any interval  $[a, b]$  with integer  $a$  and  $b$  such that  $0 < a < b$ , we have

$$\left| \sum_{a < n \leq b} \frac{1}{\ln n} - \int_a^b \frac{dt}{\ln t} \right| \leq \int_a^b \frac{dt}{t \cdot (\ln t)^2} \leq \frac{b-a}{a \cdot (\ln a)^2} \quad (10)$$

**Proof.**

Due to the *Euler's summation formula*, for positive integer numbers  $a$  and  $b$  and a function  $f$  with a continuous derivative  $f'$  on  $[a, b]$ , we have

$$\sum_{a < n \leq b} f(n) = \int_a^b f(t) dt + \int_a^b (t - [t]) f'(t) dt, \text{ where } [t] \text{ denotes an integer part of } t.$$

$$\left| \sum_{a < n \leq b} \frac{1}{\ln n} - \int_a^b \frac{dt}{\ln t} \right| = \left| \int_a^b (t - [t]) f'(t) dt \right| \leq \int_a^b |f'(t)| dt \leq \frac{b-a}{a \cdot (\ln a)^2}$$

**Q.E.D.**

Consider the *Eulerian logarithmic integral*  $Li(x) = \int_2^x \frac{dt}{\ln t}$  to evaluate  $\sum_{i=n_k+1}^{n_{k+1}} r_{ki}$ .

**Lemma 2.**

$$\sum_{i=n_k+1}^{n_{k+1}} r_{ki} = \int_{n_k}^{n_{k+1}} \frac{dt}{\ln t} = Li(n_{k+1}) - Li(n_k) + O\left(\frac{n_{k+1}}{\ln^2(n_{k+1})}\right) = \frac{n_{k+1}}{\ln n_{k+1}} - \frac{n_k}{\ln n_k} + O\left(\frac{n_{k+1}}{\ln^2 n_{k+1}}\right)$$

**Proof.**

We have

$$\text{By using approximation [7]: } li(x) = \int_0^x \frac{dt}{\ln t} = \frac{x}{\ln x} + O\left(\frac{1}{\ln^2 x}\right),$$

$$\text{we have: } Li(x) = li(x) - li(2), \text{ where } li(x) = \int_0^x \frac{dt}{\ln t} = \frac{x}{\ln x} + O\left(\frac{x}{\ln^2 x}\right).$$

$$\text{This implies } Li(n_{k+1}) - Li(n_k) = \frac{n_{k+1}}{\ln n_{k+1}} - \frac{n_k}{\ln n_k} + O\left(\frac{n_{k+1}}{\ln^2 n_{k+1}}\right).$$

**Q.E.D.**

Consider now a diffusion process  $\hat{\pi}(t)$  given by stochastic integral:

$$\hat{\pi}(t) = \int_0^t m(s) ds + \int_0^t \sigma(s) dw(s) \tag{11}$$

where  $m(t) = \frac{c}{\ln(\tau^{-1}(t))}$ ,  $\sigma(t) = \frac{1}{2} \cdot m(t) \cdot (1 - m(t))$ ,  $c = \frac{2}{e^\gamma}$ ,  $0 \leq t \leq 1$ ;  $\tau^{-1}(t) = x$ .

with the transition probability  $u(t, x, A) = P\{\hat{\pi}(t) \in A \mid \hat{\pi}(t_0) = 0\}$ .

Here  $w(t)$  is a process of Brownian motion on  $0 \leq t \leq 1$ .

The semigroup of linear operators  $U_t$  is defined on the space of bounded measurable functions by  $(U_t f)(x) = \int f(y)u(t, x, dy)$ .

We have the *infinitesimal generator of the semigroup*  $U_t$  given by the formula:

$$(Lf)(x) = \lim_{\Delta t \rightarrow 0} \frac{(U_{t+\Delta t} f)(x) - f(x)}{\Delta t}.$$

On the set of twice continuously differentiable functions  $C^2(\mathbb{R})$  the generator  $L$

takes a form of a differential operator  $(Lf)(x) = m(x) \frac{\partial f}{\partial x} + \frac{1}{2} \sigma^2(x) \frac{\partial^2 f}{\partial x^2}$ .

The function  $V(t, x) = (U_t f)(x) = E[Y(t) \mid Y(t_0) = x] = \int f(y)u(t, x, dy)$  satisfies the equation

$$\frac{\partial V}{\partial t} = \frac{1}{2} \sigma^2(x) \frac{\partial^2 V}{\partial x^2} + m(x) \frac{\partial V}{\partial x} \quad (12)$$

with the initial condition  $V(t_0, x) = f(x)$ . Taking as an initial condition  $\delta$ -function, we have  $V(t, x) = u(t, x, y)$ , called a *fundamental solution* to (12).

This means that the transition probability has a density  $u(t, x, y)$ , so that

$$P\{Y(t) \in A \mid Y_0(t) = x\} = \int_A u(t, x, y) dy.$$

By applying the generalized limit theorem [22, 23] about convergence of random walks  $\pi^\Delta(t_k)$  as  $|\Delta| = \max_{1 \leq k \leq K} \Delta t_k \rightarrow 0$ ,  $N_{\min} \rightarrow \infty$  to diffusion processes (6.9), we obtain an approximation of  $\{\pi(n, \omega)\}_{n \in \mathbb{W}}$  in terms of diffusion processes, defined for expanding intervals  $[N_{\min}, N_{\max}]$  of approximation on  $\mathbb{N}$ .



**Theorem 1.**

Transition probabilities

$P\{\pi^\Delta(t_{k+1}) \in E \mid \pi^\Delta(t_k) = x_k, \pi^\Delta(t_{k-1}) = x_{k-1}, \dots, \pi^\Delta(t_0) = x_0\}$ , where  $\vec{x}_k = (x_1, \dots, x_k) \in \mathbb{N}^k$ ,

of the defined above non-Markov random walks  $\{\pi^\Delta(t_k) \mid N_{\min} < k < N_{\max}\}$  converge weakly to the transition probabilities of the diffusion process  $\hat{\pi}(t)$  given by the stochastic integral

$$\hat{\pi}(t) = \int_0^t \hat{m}(s) ds + \int_0^t \hat{\sigma}(s) dw(s), \quad (13)$$

where  $\hat{m}(t) = \frac{c}{\ln(\tau^{-1}(t))}$ ,  $\hat{\sigma}(t) = \frac{1}{2} \cdot \hat{m}(t) \cdot (1 - \hat{m}(t))$ ,  $c = \frac{2}{e^\gamma}$ ,  $0 \leq t \leq 1$ ,

$\tau^{-1}(t) = x$ ,  $N_{\min} \leq n \leq N_{\max}$ ,  $\tau(N_{\min}) = 0$ ,  $\tau(N_{\max}) = 1$ ,  $c = \frac{2}{e^\gamma} \approx 1.122918968$

with the Euler's constant  $\gamma = \sum_{m \leq n} \frac{1}{m} - \ln n + O\left(\frac{1}{n}\right)$ ,  $\gamma \approx 0.577215664$ ,

as  $|\Delta| = \max_{1 \leq k \leq K} \Delta t_k \rightarrow 0$ ,  $N_{\min} \rightarrow \infty$ .

**Proof.**

Since  $\sum_{k=1}^K \Delta t_k = 1$ , to due to Lemma 3.1.1, we have:

$$\sum_{k=1}^K \frac{1}{\ln n_k} \leq \frac{K}{\ln(N_{\min})} \cdot \left(1 + O\left(\frac{1}{\ln(N_{\min})}\right)\right) \rightarrow 0, \text{ while } \frac{K}{N_{\max}} \rightarrow 0$$

Then, formulas (3.1.7), due to the second Merten's theorem (the *Merten's formula*) [3, 19], imply:

$$\begin{aligned}
& \sum_{k=1}^K \left[ |m^\Delta(t_k, \bar{x}_k) - m(t_k)| + |(\sigma^\Delta(t_k, \bar{x}_k))^2 - (\sigma(t_k, \bar{x}_k))^2| \right] \cdot \Delta t_k \\
&= 2 \cdot \sum_{k=1}^K \left[ \left| \frac{1}{\ln(n_k)} - \frac{1}{\ln(n_k)} \cdot \Delta t_k + \frac{1}{\ln(n_k)} \cdot O\left(\frac{1}{\ln(n_k)}\right) \right| \right] \leq 2 \cdot \sum_{k=1}^K \left[ \frac{1}{\ln(n_k)} \left( 1 - \Delta t_k + O\left(\frac{1}{\ln(n_k)}\right) \right) \right] \\
&\leq 2 \cdot \max_{1 \leq k \leq K} \left| 1 - \Delta t_k + O\left(\frac{1}{\ln(n_k)}\right) \right| \cdot \sum_{k=1}^K \frac{1}{\ln n_k} \rightarrow 0, \text{ as } N_{\min} \rightarrow \infty
\end{aligned}$$

For  $g(k) = \pi(k+1) - \pi(k) \leq 1$ , we have  $P\{\eta^\Delta(k) > 1\} = 0$  for all  $k$ , so that all conditions are satisfied to apply the limit theorems for random walks proved in [22, 23, 24].

**Q.E.D.**

The figures below illustrate graphically the diffusion approximation of distribution of primes in terms of  $\pi(n)$  on different intervals of the argument.

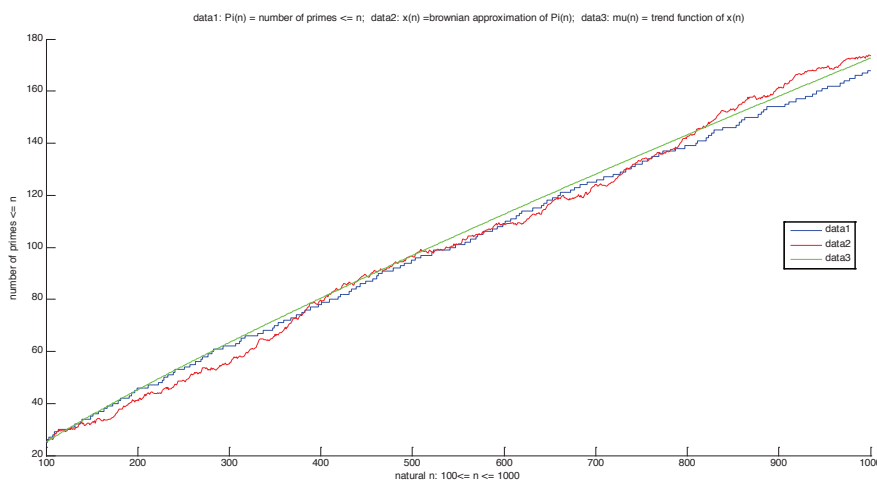
Legend for the graphs on the following figures:

data1:  $\pi(n)$  = exact number of primes  $\leq n$

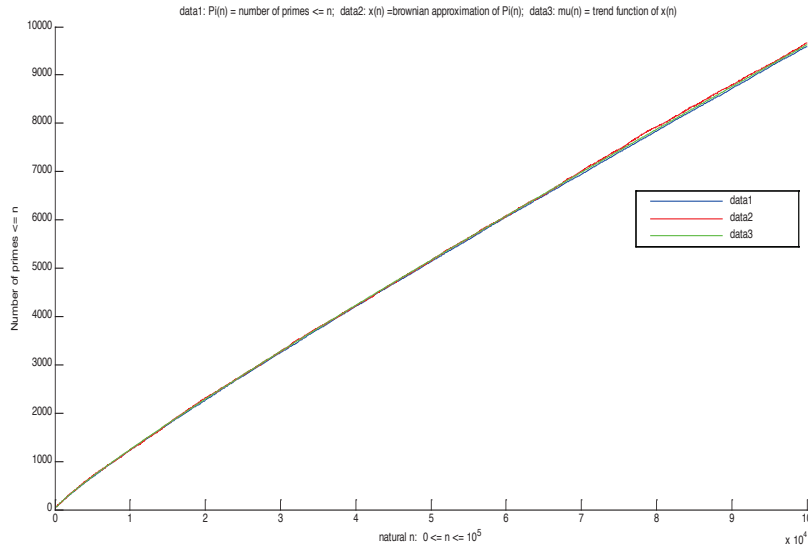
data2: Brownian approximation  $X_n = \mu(n) + \xi_n \cdot \sigma(n)$  of  $\pi(n)$

data3: Trend function  $\mu(n)$  of  $X_n$

Approximation of  $\pi(n)$  for  $n: 100 \leq n \leq 1000$



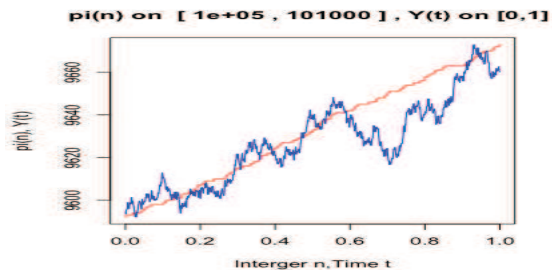
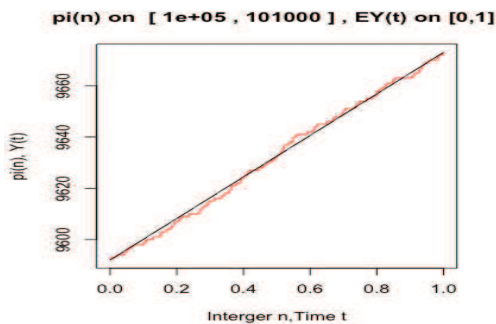
## Approximation of $\pi(n)$ for $n: 0 \leq n \leq 10^5$



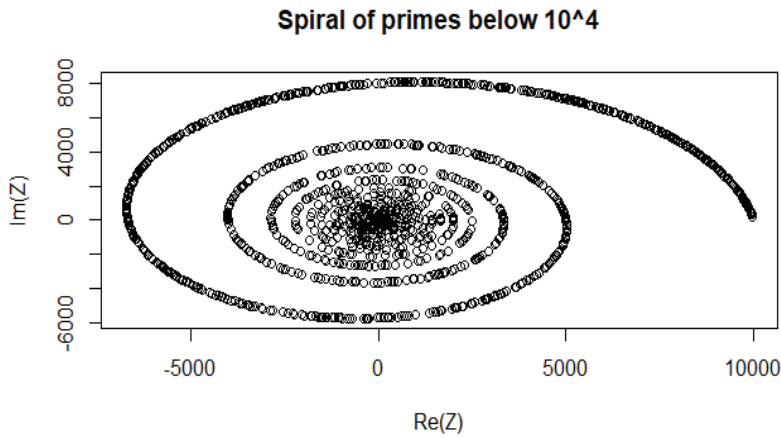
## Approximation of $\pi(n)$ for $n: 5 \times 10^5 \leq n \leq 10^6$

On figures below there are the graphs of paths described evolution of the ‘walk’ of a counts  $\{\pi(n) | n \in \mathbb{N}\}$  of consecutive primes restricted to the intervals  $[N_{\min}, N_{\max}] \subset \mathbb{N}$  and approximating diffusion processes :

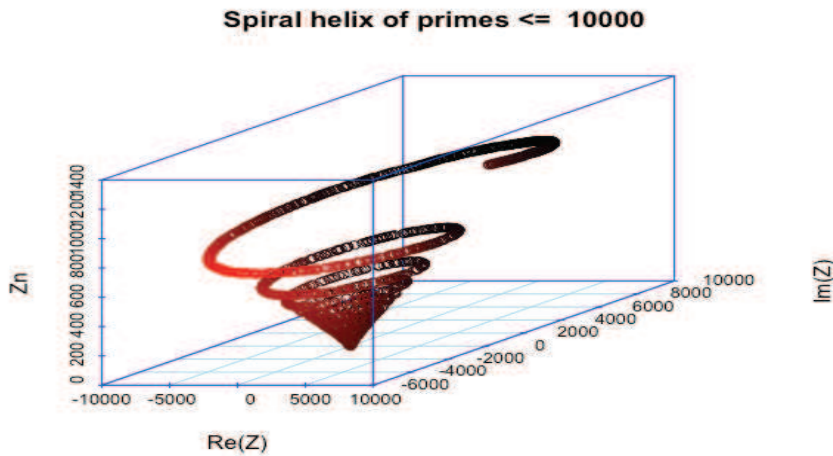
$$Y(t) = \hat{\pi}^\Delta(t) \text{ and their expectations } EY(t) \text{ for } t \in [0,1] \subset \mathbb{R}.$$



The sequence of vectors  $(\bar{p}(n), \bar{r}(n))$ ,  $(n = 2, 3, \dots)$  created by consecutive  $n$  primes and the residual values  $\bar{r} = \text{mod}(n, \bar{p})$ , allows an interesting 3D presentation. In each pair  $(\bar{p}(n), \bar{r}(n))$  vector of primes  $\bar{p}(n)$  represents a ‘radial’ component, while the vector of residuals  $\bar{r}(n)$ , due to its natural periodicity, represents a ‘circular’ component.



Denote  $z_k = p_k \cdot \exp\left(2\pi i \cdot \frac{r_k}{p_k - 1}\right)$ ,  $r_k = \text{mod}(n, p_k)$ ,  $(k = 1, 2, 3, \dots)$  - a sequence of complex numbers and the vector  $\vec{z}(n) = (z_1, \dots, z_n)$ . Then for any  $n > 2$  vector  $\vec{z}(n)$  takes a shape of a spiral helix as in the pictures below.



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